Approximate Bayesian Computation in Insurance

Patrick J. Laub and Pierre-Olivier Goffard





Motivation

Have a random number of claims $N \sim p_N(\,\cdot\,; \boldsymbol{\theta}_{\mathrm{freq}})$ and the claim sizes $U_1, \ldots, U_N \sim f_U(\,\cdot\,; \boldsymbol{\theta}_{\mathrm{sev}})$.

We aggregate them somehow, like:

- aggregate claims: $X = \sum_{i=1}^{N} U_i$
- maximum claims: $X = \max_{i=1}^N U_i$
- stop-loss: $X = (\sum_{i=1}^N U_i c)_+$.

Question: Given a sample X_1, \ldots, X_n of the summaries, what is the $\theta = (\theta_{\text{freq}}, \theta_{\text{sev}})$ which explains them?

E.g. a reinsurance contract

Likelihoods

For simple rv's we know their likelihood (normal, exponential, gamma, etc.).

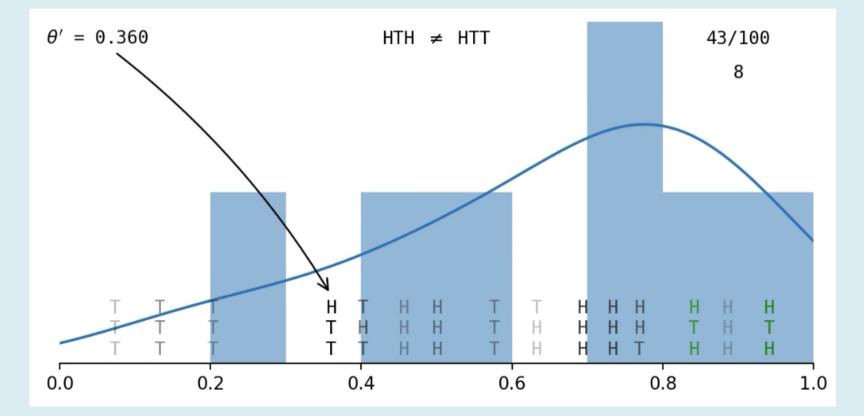
When simple rv's are combined, the resulting thing rarely has a tractable likelihood.

 $X_1, X_2 \stackrel{\mathrm{i.i.d.}}{\sim} f_X(\,\cdot\,) \Rightarrow X_1 + X_2 \sim$ Intractable likelihood!

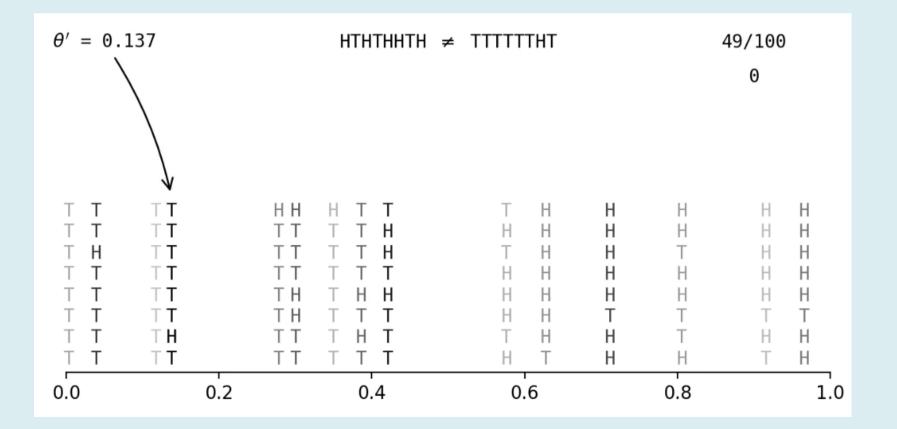
Usually it's still possible to simulate these things...

Approximate Bayesian Computation

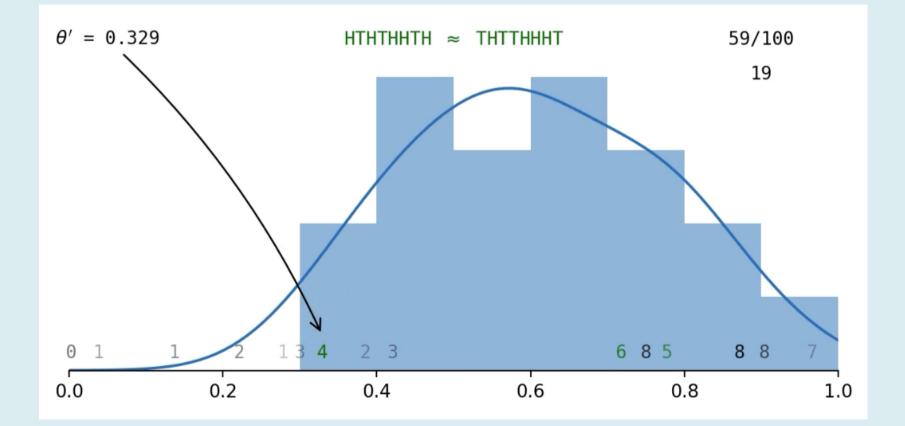
Example: Flip a coin a few times and get $(x_1, x_2, x_3) = (H, T, H)$; what is $\pi(\theta | \boldsymbol{x})$?



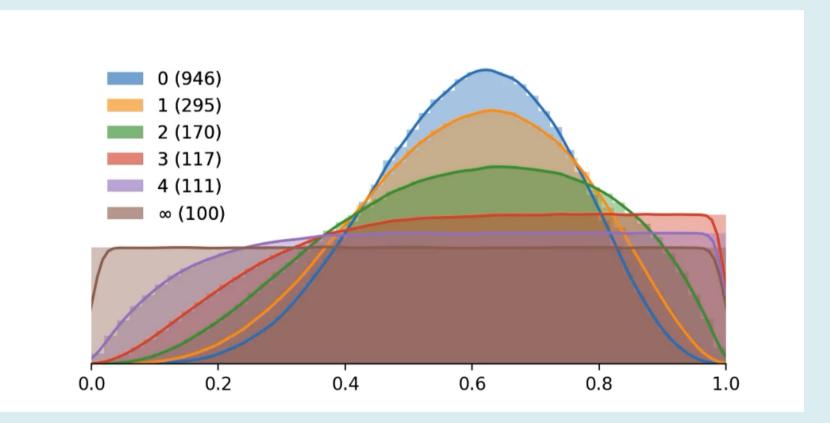
Getting an exact match of the data is hard...



Accept fake data that's close to observed data



The 'approximate' part of ABC



Does it work in theory?

We sample the *approximate / ABC posterior* $\pi_{\epsilon}(\boldsymbol{\theta} \mid \boldsymbol{x}_{\text{obs}}) \propto \pi(\boldsymbol{\theta}) \times \mathbb{P}(||\boldsymbol{x}_{\text{obs}} - \boldsymbol{x}^{*}|| \leq \epsilon \text{ where } \boldsymbol{x}^{*} \sim \boldsymbol{\theta}),$ but we care about the true posterior $\pi(\boldsymbol{\theta} \mid \boldsymbol{x}_{\text{obs}})$.

Proposition: Say we have continuous data \boldsymbol{x}_{obs} , and our prior $\pi(\boldsymbol{\theta})$ has bounded support.

If for some $\epsilon > 0$ we have

$$\sup_{(oldsymbol{z},oldsymbol{ heta}):\mathcal{D}(oldsymbol{z},oldsymbol{x}_{ ext{obs}})<\epsilon,oldsymbol{ heta}\inoldsymbol{arDelta}}\piig(oldsymbol{z}\midoldsymbol{ heta}ig)<\infty$$

then for each $oldsymbol{ heta}\inoldsymbol{\Theta}$

$$\lim_{\epsilon o 0} \pi_\epsilon(oldsymbol{ heta} \mid oldsymbol{x}_{ ext{obs}}) = \; \pi(oldsymbol{ heta} \mid oldsymbol{x}_{ ext{obs}}) \; .$$

Rubio and Johansen (2013), *A simple approach to maximum intractable likelihood estimation,* Electronic Journal of Statistics.

Mixed data

We filled in some of the blanks for *mixed data*. Our data was mostly continuous data but had an atom at 0.

Get

$$\lim_{\epsilon o 0} \pi_\epsilon(oldsymbol{ heta} \mid oldsymbol{x}_{ ext{obs}}) = \; \pi(oldsymbol{ heta} \mid oldsymbol{x}_{ ext{obs}})$$

when

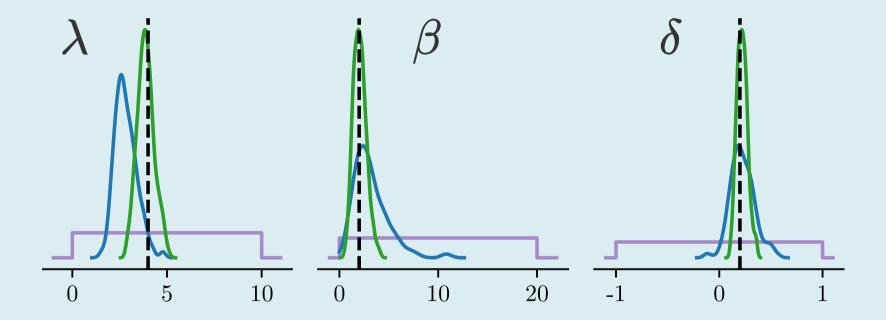
$$\mathcal{D}(oldsymbol{z},oldsymbol{x}_{ ext{obs}}) = egin{cases} \mathcal{D}^+(oldsymbol{z}^+,oldsymbol{x}_{ ext{obs}^+}) & ext{if}\, \# ext{Zeros}(oldsymbol{z}) = \# ext{Zeros}(oldsymbol{x}_{ ext{obs}}), \ \infty & ext{otherwise}. \end{cases}$$

Does it work in practice?

In other words, when does it break & how slow is it?

Dependent example

- frequency $N \sim \mathsf{Poisson}(\lambda = 4)$,
- severity $U_i \mid N \sim \mathsf{DepExp}(\beta = 2, \delta = 0.2)$, defined as $U_i \mid N \sim \mathsf{Exp}(\beta \times e^{\delta N})$,
- summation summary $X = \sum_{i=1}^{N} U_i$



ABC posteriors based on 50 X's and on 250 X's given uniform priors.

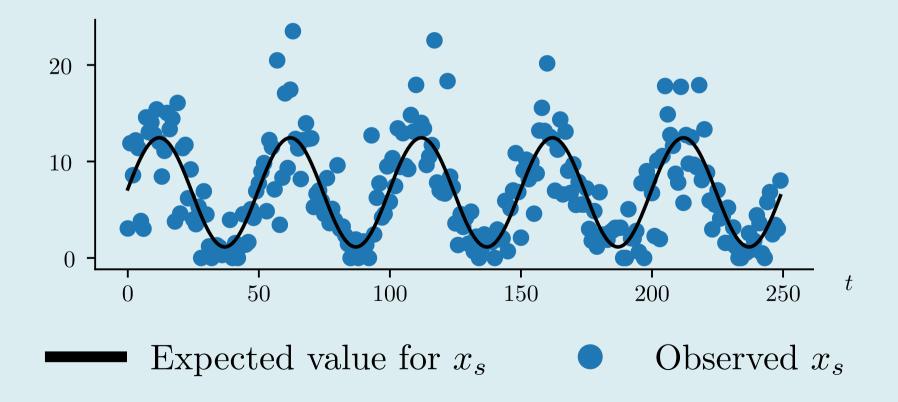
J. Garrido, C. Genest, and J. Schulz (2016), *Generalized linear models for dependent frequency and severity of insurance claims*, IME.

pip install approxbayescomp

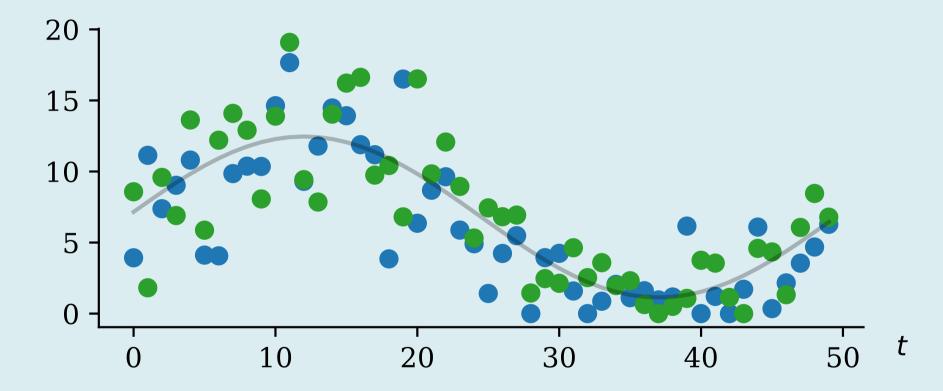
```
import approxbayescomp as abc
 1
 2
  # Load data to fit
 3
  obsData = ...
 4
 5
  # Frequency-Loss Model
 6
   freq = "poisson"
 7
   sev = "frequency dependent exponential"
 8
  psi = abc.Psi("sum") # Aggregation process
 9
10
11 # Fit the model to the data using ABC
   prior = abc.IndependentUniformPrior([(0, 10), (0, 20), (-1, 1)])
12
13 model = abc.Model(freq, sev, psi, prior)
14 fit = abc.smc(numIters, popSize, obsData, model)
```

Time-varying example

- Claims form a Poisson process point with arrival rate $\lambda(t) = a + b[1 + \sin(2\pi ct)]$.
- The observations are $X_s = \sum_{i=N_{s-1}}^{N_s} U_i$.
- Frequencies are CPoisson $(a = 1, b = 5, c = \frac{1}{50})$ and sizes are $U_i \sim \text{Lognormal}(\mu = 0, \sigma = 0.5)$.

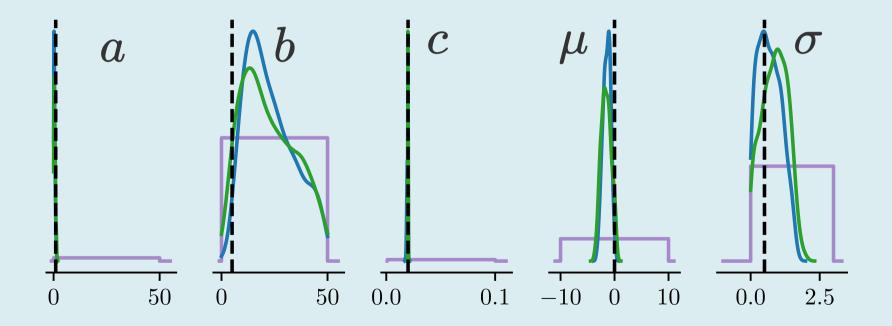


From the same heta the data is **random**



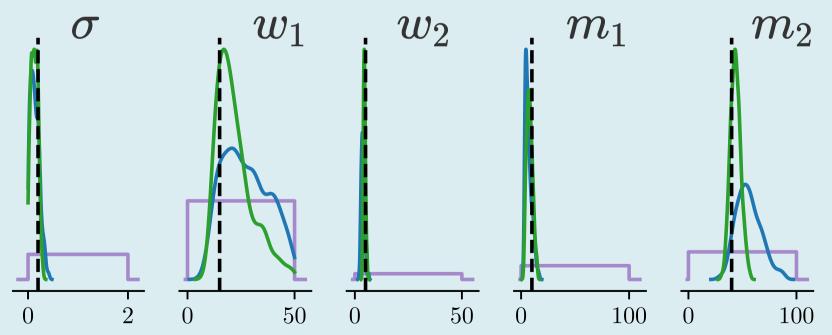
Time-varying example

- Claims form a Poisson process point with arrival rate $\lambda(t) = a + b[1 + \sin(2\pi ct)]$.
- The observations are $X_s = \sum_{i=N_{s-1}}^{N_s} U_i$.
- Frequencies are CPoisson $(a = 1, b = 5, c = \frac{1}{50})$ and sizes are $U_i \sim \text{Lognormal}(\mu = 0, \sigma = 0.5)$.
- ABC posteriors based on 50 *X*'s and on 250 *X*'s with uniform priors.



Bivariate example

- Two lines of business with dependence in the claim frequencies.
- Say $\Lambda_i \sim \mathsf{Lognormal}(\mu \equiv 0, \sigma = 0.2).$
- Claim frequencies $N_i \sim \mathsf{Poisson}(\Lambda_i w_1)$ and $M_i \sim \mathsf{Poisson}(\Lambda_i w_2)$ for $w_1 = 15$, $w_2 = 5$.
- Claim sizes for each line are $Exp(m_1 = 10)$ and $Exp(m_2 = 40)$.
- ABC posteriors based on 50 *X*'s and on 250 *X*'s with uniform priors.



Streftaris and Worton (2008), *Efficient and accurate approximate Bayesian inference with an application to insurance data*, Computational Statistics & Data Analysis

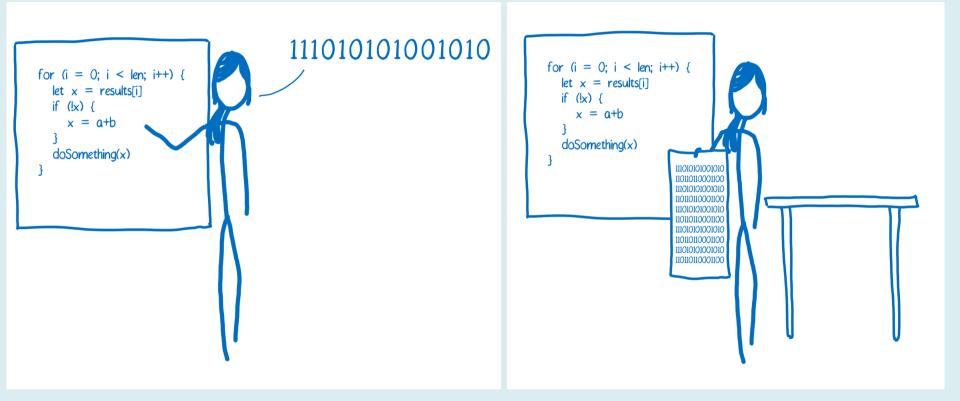
ABC turns a statistics problem into a programming problem

... to be used as a last resort

Compiled code

(use *numba*)

Interpreters & compilers



C.f. Lin Clark (2017), A crash course in JIT compilers

Interpreted version

```
1
   def sample geometric exponential sums(T, p, \mu):
 2
     X = np.zeros(T)
 3
 4
    N = rnd.geometric(p, size=T)
 5
     U = rnd.exponential(\mu, size=N.sum())
 6
 7
   i = 0
 8
     for t in range(T):
 9
       X[t] = U[i:i+N[t]].sum()
10
       i += N[t]
11
12
   return X
```

Compiled by numba

```
from numba import njit
 1
 2
 3
   @njit()
   def sample geometric_exponential_sums(T, p, µ):
 4
 5
     X = np.zeros(T)
 6
 7
     N = rnd.geometric(p, size=T)
     U = rnd.exponential(µ, size=N.sum())
8
 9
   i = 0
10
11
   for t in range(T):
12
       X[t] = U[i:i+N[t]].sum()
13
       i += N[t]
14
15
    return X
```

First run is compiling (500 ms), but after we are down from 2.7 ms to 164 µs (16x speedup)

Original code: 1.7 s Basic profiling with **snakeviz**: 5.5 ms, 310x speedup + Vectorisation/preallocation with **numpy**: 2.7 ms, 630x speedup + Compilation with **numba**: 164 µs, 10,000x speedup

And potentially:

+ Parallel over 80 cores: say another 50x improvement, so overall 50,000x speedup.

Take-home messages

- What is ABC?
- ABC turns a *statistics problem* into a *programming problem*
- pip install approxbayescomp